

A multiplicative-regression model to compare the effect of factors associated with the time to graft failure between first and second renal transplant

Katy Trébern-Launay, Magali Giral, Yohann Foucher

EA 4275 Biostatistics, Clinical Research and Subjective Measures in Health Sciences,
Nantes University, Nantes
Transplantation, Urology and Nephrology Institute (ITUN), Nantes Hospital and
University, Inserm U1064, Centaure, Nantes, France

August 20, 2012

Introduction

Kidney transplantation

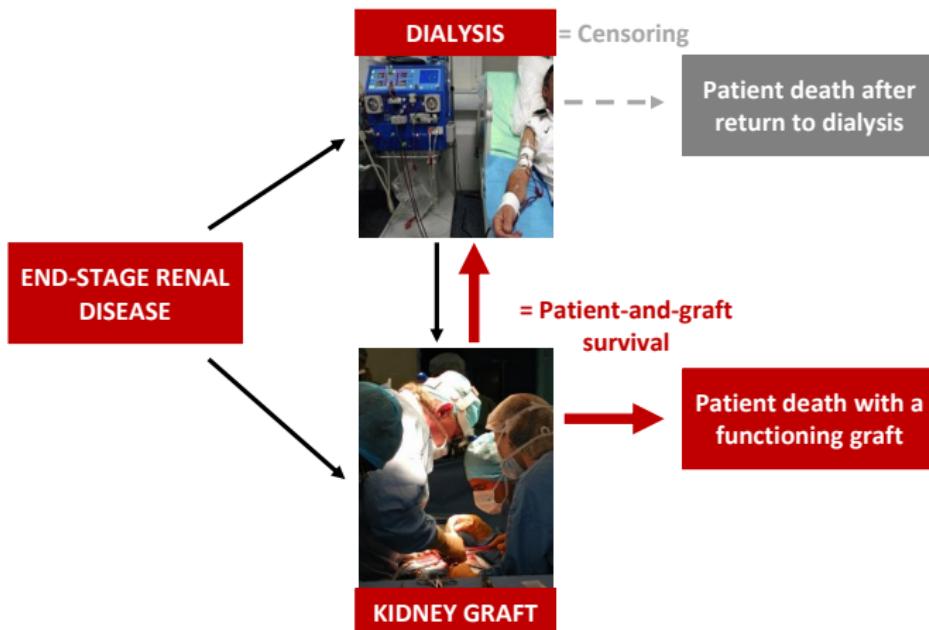
Introduction

Materials and Methods

Results

Conclusion

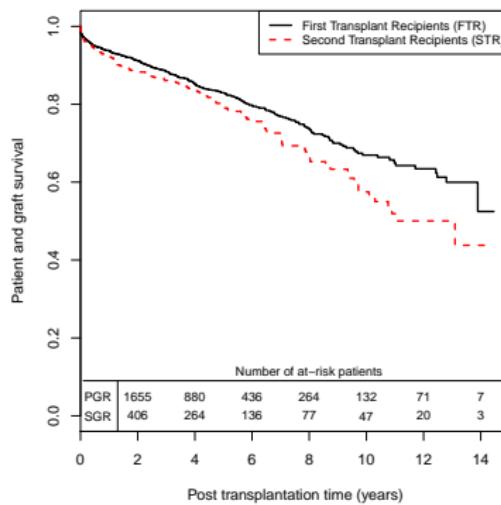
Definition of the graft failure



Introduction

Objectives

Close patient-and-graft survival between first and second graft



Objective

- Are risk factors associated with graft failure comparable between first and second grafts ?

Introduction

Materials and Methods

Results

Conclusion

Limits of classical survival models

- Test of interaction between each covariate and graft rank
- Only covariates common to first and second grafts

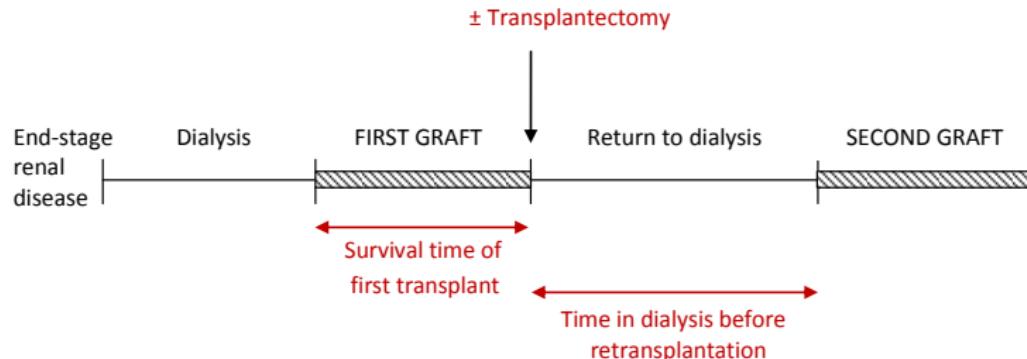


FIGURE 1: Clinical trajectory before second graft.

Classical approach

- Additive-regression model for relative survival
(Estève et al. Stat in Med 1990)
- Endpoint = mortality related to chronic diseases
- The expected mortality is based on general population

Proposed approach

- Multiplicative-regression model for relative survival
(Andersen et al. Stat in Med 1989)
- Endpoint = **graft failure** (return to dialysis or patient death)
- The expected graft failure hazard is **estimated** in a control group (first graft)

Materials

Introduction

Materials and Methods

Results

Conclusion

Inclusion criteria

French DIVAT database

- Centers : Nantes, Necker, Nancy, Toulouse, Montpellier, Lyon
- Adult recipients
- Transplanted from 1996 to 2010
- Under mycophenolate mofetil and steroids at transplantation



Group of interest

566 second transplant
recipients (**STR**)

Control group

2206 first transplant
recipients (**FTR**)

Multiplicative-regression models for relative survival

- The hazard function

$$h^{(o)}(t_i, z_i) = h^{(e)}(t_i, z_i^{(e)}) \cdot h^{(r)}(t_i, z_i^{(r)})$$



Observed hazard function in the STR group

z_i = covariates associated with the observed hazard

Expected hazard function in the FTR group

$z_i^{(e)}$ = subset of z_i , associated with the expected hazard

Relative hazard function in the STR group

$z_i^{(r)}$ = subset of z_i , associated with the relative hazard

STEP 1

STEP 2

STEP 1 :

Estimation of the expected hazard function ($N^{(e)} = 2206$ FTR)

- Parametric model and proportional hazards assumption

$$h^{(e)}(t_i, z_i^{(e)}) = h_0^{(e)}(t_i) \exp\left(\sum_{j=1}^{p^{(e)}} \beta_j^{(e)} z_{i,j}^{(e)}\right)$$

- $h_0^{(e)}(t_i)$ is a piecewise function
- Maximum-likelihood estimation

$$\log \mathcal{L} = \sum_{i=1}^{N^{(e)}} \{\delta_i \log(h^{(e)}(t_i, z_i^{(e)})) - H^{(e)}(t_i, z_i^{(e)})\}$$

with $\delta_i = 1$ if the graft failure is observed
 $\delta_i = 0$ if the event is right-censored

STEP 2 :

Estimation of the relative hazard function ($N^{(r)} = 566$ STR)

- Parametric model and proportional hazards assumption

$$h^{(r)}(t_i, z_i^{(r)}) = h_0^{(r)}(t_i) \exp\left(\sum_{j=1}^{p^{(r)}} \beta_j^{(r)} z_{i,j}^{(r)}\right)$$

- $h_0^{(r)}(t_i)$ is a piecewise function
- Maximum-likelihood estimation

$$\log \mathcal{L} = \sum_{i=1}^{N^{(o)}} \{\delta_i \log(h^{(o)}(t_i, z_i)) - H^{(o)}(t_i, z_i)\}$$



$$h^{(e)}(t_i, z_i^{(e)}) \quad h^{(r)}(t_i, z_i^{(r)}) \quad \int_0^{t_i} h^{(e)}(u, z_i^{(e)}) \quad h^{(r)}(u, z_i^{(r)}) du$$

Methods

Interpretation of the regression coefficient

CASE 1 :

For $z_1^{(r)} \notin z_j^{(e)}$ $\Rightarrow \exp(\beta_1^{(r)}) = \text{hazard ratio}$

$$\text{HR}_{z_1=1/z_1=0}^{(o)} = \frac{h^{(e)}(t_i, z_i^{(e)}) \ h_0^{(r)}(t_i) \ exp(\sum_{j=1}^{p(r)} \beta_j^{(r)} z_{i,j}^{(r)})}{h^{(e)}(t_i, z_i^{(e)}) \ h_0^{(r)}(t_i) \ exp(\sum_{j=2}^{p(r)} \beta_j^{(r)} z_{i,j}^{(r)})} = \exp(\beta_1^{(r)})$$

CASE 2 :

For $z_1^{(r)} \in z_j^{(e)}$ $\Rightarrow \exp(\beta_1^{(r)}) = \text{weighting factor of HR}$

$$\begin{aligned} \text{HR}_{z_1=1/z_1=0}^{(o)} &= \frac{h_0^{(e)}(t_i) \ exp(\sum_{j=1}^{p(e)} \beta_j^{(e)} z_{i,j}^{(e)}) \ h_0^{(r)}(t_i) \ exp(\sum_{j=1}^{p(r)} \beta_j^{(r)} z_{i,j}^{(r)})}{h_0^{(e)}(t_i) \ exp(\sum_{j=2}^{p(e)} \beta_j^{(e)} z_{i,j}^{(e)}) \ h_0^{(r)}(t_i) \ exp(\sum_{j=2}^{p(r)} \beta_j^{(r)} z_{i,j}^{(r)})} \\ &= \exp(\beta_1^{(e)}) \ exp(\beta_1^{(r)}) \end{aligned}$$

Introduction

Materials and Methods

Results

Conclusion

Methods

Covariates tested in the relative hazard

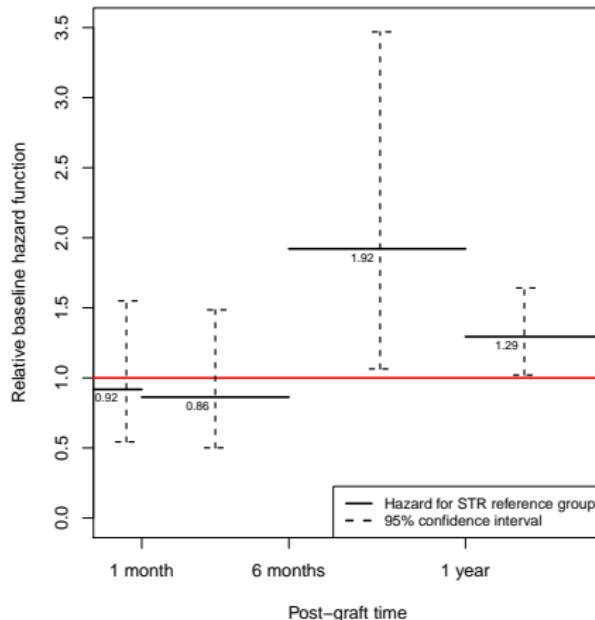
Introduction
Materials and Methods
Results
Conclusion

Covariates	Included in $z^{(e)}$
Recipient age (≥ 55 years / < 55 years)	✓
Recipient gender (male / female)	✓
Causal nephropathy (recurrent / non recurrent)	✓
History of comorbidities (positive / negative)	✓
Body mass index ($\geq 30 \text{ kg.m}^{-2}$ / $< 30 \text{ kg.m}^{-2}$)	✓
Anti-class I or II PRA (positive / negative)	✓
Dialysis prior transplantation (positive / negative)	
Recipient EBV or CMV serology (positive / négative)	
Type of donor (deceased donor / living donor)	✓
Donor age (≥ 55 years / < 55 years)	✓
Donor EBV serology (positive / négative)	✓
Donor gender (male / female)	
Cause of donor death (cerebro-vascular / other)	
Donor serum creatinine ($\geq 133 \mu\text{mol/l}$ / $< 133 \mu\text{mol/l}$)	
Donor CMV serology (positive / négative)	
Transplantation period (< 2005 / ≥ 2005)	✓
Number of HLA-A-B-DR mismatches (> 4 / ≤ 4)	✓
Induction therapy (depleting / non depleting)	✓
Cold ischemia time ($\geq 24\text{h}$ / $< 24\text{h}$)	✓
Survival time of the first transplant (< 1 year / ≥ 1 year)	
Time before retransplantation (> 3 years / ≤ 3 years)	
First graft transplantectomy (positive / négative)	

Results

Relative baseline hazard function $\Leftrightarrow \mathbf{z}^{(r)} = \mathbf{0}$

Introduction
Materials and Methods
Results
Conclusion



Results

Hazard ratio obtained from the multivariate model

Introduction

Materials and Methods

Results

Conclusion

Final multivariate model

Covariates	Expected (FTR)	Relative (STR)	Observed (STR)	95% CI	p value
Transplant period (< 2005/ \geq 2005)	1.37	1.27	-	0.83 - 1.95	0.2604
Recipient gender (male / female)	1.19	0.68	-	0.45 - 1.02	0.0645
Recipient age (\geq 55 years/<55 years)	1.55	1.61	-	1.03 - 2.52	0.0387
Donor age (\geq 55 years/<55 years)	1.37	0.59	-	0.37 - 0.95	0.0294
Type of donor (deceased/living)	2.91	0.33	-	0.12 - 0.91	0.0332
Donor gender (male / female)	-	-	1.57	1.01 - 2.45	0.0443
Retransplant time (>3 years/ \leq 3 years)	-	-	2.06	1.33 - 3.20	0.0012

xxx = forced covariates

Results

Hazard ratio obtained from the multivariate model

Introduction
Materials and Methods
Results
Conclusion

CASE 1 : $z_1^{(r)} \notin z_j^{(e)}$

Covariates	Expected (FTR)	Relative (STR)	Observed (STR)	95% CI	p value
Transplant period (< 2005/ \geq 2005)	1.37	1.27	-	0.83 - 1.95	0.2604
Recipient gender (male / female)	1.19	0.68	-	0.45 - 1.02	0.0645
Recipient age (\geq 55 years/<55 years)	1.55	1.61	-	1.03 - 2.52	0.0387
Donor age (\geq 55 years/<55 years)	1.37	0.59	-	0.37 - 0.95	0.0294
Type of donor (deceased/living)	2.91	0.33	-	0.12 - 0.91	0.0332
Donor gender (male / female)	-	-	1.57	1.01 - 2.45	0.0443
Retransplant time (>3 years/ \leq 3 years)	-	-	2.06	1.33 - 3.20	0.0012

xxx = forced covariates



$$HR^{(o)} = \exp(\beta_1^{(r)})$$

Results

Hazard ratio obtained from the multivariate model

Introduction

Materials and Methods

Results

Conclusion

CASE 2 : $z_1^{(r)} \in z_j^{(e)}$

Covariates	Expected (FTR)	Relative (STR)	Observed (STR)	95% CI	p value
Transplant period (< 2005/≥2005)	1.37	1.27	-	0.83 - 1.95	0.2604
Recipient gender (male / female)	1.19	0.68	-	0.45 - 1.02	0.0645
Recipient age (≥55 years/<55 years)	1.55	1.61	2.50	1.03 - 2.52	0.0387
Donor age (≥55 years/<55 years)	1.37	0.59	0.81	0.37 - 0.95	0.0294
Type of donor (deceased/living)	2.91	0.33	0.96	0.12 - 0.91	0.0332
Donor gender (male / female)	-	-	1.57	1.01 - 2.45	0.0443
Retransplant time (>3 years/≤3 years)	-	-	2.06	1.33 - 3.20	0.0012

xxx = forced covariates



$$\text{HR}^{(o)} = \exp(\beta_1^{(e)}) \exp(\beta_1^{(r)})$$

Conclusion

Introduction

Materials and Methods

Results

Conclusion

Clinical conclusions

- A particular attention to recipient age for clinical practice when faced a second transplantation should be paid
- A selection bias ?
Only transplants from "good quality" donors are proposed for STR when the donor is aged or deceased.
- An early effect of immunisation ?
The immunisation might take over the effect of other factors (donor age and donor type) for STR.

Statistical conclusion

- No necessity to test interactions between covariates and graft rank
- Possibility to take into account specific covariates for interest groups

Limits

- Parameters of the expected hazard function (STEP 1) were afterwards considered as constant when used for the estimation of the relative hazard (STEP 2)
- Proportional hazard model with a piecewise baseline function were chosen for the estimation of both baseline hazards functions
- Only time-invariant covariates were included in the model

- ⇒ A Monte-Carlo approach is in process
- ⇒ A more flexible model for instance with spline functions
- ⇒ A generalisation with time-dependent covariates

Introduction

Materials and Methods

Results

Conclusion

Particular thanks to :

- The EA-4275 Unit
- The DIVAT-Biostatistics Unit :
 - Magali Giral, Scientific council coordinator
 - Pascal Daguin, Computer engineer
 - Sandra Lefloch, Clinical research associate
 - Yohann Foucher, Associate professor in Biostatistics
 - Etienne Dantan, Associate professor in Biostatistics
 - Marine Lorent, PhD Student in Biostatistics
 - Florence Gilliaizeau, PhD Student in Biostatistics
 - Philippe Tessier, Health Economist
 - Florent Leborgne, Trainee in Biostatistics

Of note

- This presentation will be available online : <http://www.divat.fr/en>